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NAVAL POSTGRADUATE SCHOOL Monterey, California



THESIS

Implementation of a Reliability Shorthand on the TI-59 Handheld Calculator

by

Hans-Eberhard Peters

October 1982

Thesis Advisor:

J.D. Esary

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Assuming constant failure rates, basic structures are used to show how the shorthand can be applied. Several examples are worked out that show, how, with component failure rates as input, a handheld calculator can be



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Two TI-59 programs are provided as a computational aid.

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Implementation of a Reliability Shorthand on the TI-59 Handheld Calculator

by

Hans-Eberhard Peters
Major, German Air Force
Dipl.-Betriebsw., Fachhochschule des Heeres I, 1974

Submitted in partial fulfillment of the requirements for the degree of

MASTER OF SCIENCE IN OPERATIONS RESEARCH

from the NAVAL POSTGRADUATE SCHOOL October, 1982



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It is shown how a reliability shorthand can be implemented on a handheld calculator.

Assuming constant failure rates, basic structures are used to show how the shorthand can be applied. Several examples are worked out that show, how, with component failure rates as input, a handheld calculator can be used to compute the reliability of a system.

Two TI-59 programs are provided as a computational aid.



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I. INTRODUCTION

Systems and components can be in either of two states: either they are functioning or they have failed. The ability, that a system stays functioning over a predetermined time interval is called its reliability. It is generally not realistic to assume that a system, say a lightbulb, will fail at a specified time, but rather that T, the time to failure, is a random variable which has a probability distribution that can be specified. The probability distribution for a time to failure is called its life distribution. In this paper we will solely be concerned with one specific type of life distribution which is especially important in reliability theory and practice, the exponential distribution. It has the property that the remaining life of a used component is independent of its age (the "memoryless" property), i.e. a functioning component is always as good as new, the failure rate is constant. The memoryless property is the basis for a reliability shorthand, one that can be implemented on a handheld calculator.

Depending on the size, structure and life distribution of a system, probability statements about its time to



failure are in general not easily achieved. Forming the sum of independent life lengths (i.e. convolving the corresponding life distributions) requires knowledge of integral calculus and computations can become rather tedious.

In the case of the exponential distribution, though, computations can be simplified by translating the problem into a simple shorthand notation and using this shorthand as input for some computing device.

In this paper we will show how a reliability shorthand can be implemented on a handheld calculator. Basic structures are used to show how the shorthand can be applied. Two TI-59 programs are provided as a computational aid. Formulas for the convolution of up to four exponential random variables can be found in Appendix A. Appendix B contains a user guide to the TI-59 programs.



II. THE CONCEPT OF A RELIABILITY SHORTHAND

A. BASIC NOTATION

The survival function of a life length can be derived from the distribution function.

Let

T : life length

 $F(t) = P(T \le t)$ be the distribution function of T

Then

$$\vec{F}(t) = P(T>t)$$

$$= 1-F(t)$$

is the survival function of T.

In the case of the exponential distribution, $\bar{F}(t) = e^{-\lambda t}$, where λ is the failure rate. Translated into shorthand, the life distribution is denoted

B. CONVOLUTION OF DISTRIBUTIONS

When independent random lives are summed up, the corresponding life distributions have to be convolved to determine the probability that the sum of the lives will exceed a specified time t. Let

T, , T2 : independent life lengths



 \vec{F}_{1} (t), \vec{F}_{2} (t) : the corresponding survival functions

 $f_1(t), f_2(t)$: the corresponding density functions

 $T = T_4 + T_2$: the total life length

Then

$$\widetilde{F}(t) = P(T>t)$$

$$= P(T_1 + T_2 > t)$$

$$= \widetilde{F}_1(t) + \int_{0}^{t} \widetilde{F}_2(t-s) f_1(s) ds.$$

This means that T will exceed a specified time t when

-either T, exceeds t

-or T, is smaller than t, say equal to s, and T2 exceeds t-s.

Integration with respect to s (i.e. summing over all possible values of s) is called the convolution of T, and T2. When T4 and T2 are both exponentially distributed with failure rates λ_4 and λ_2 , i.e.

$$\overline{F}_{1}(t) = e^{-\lambda_{1}t}$$
 $\overline{F}_{2}(t) = e^{-\lambda_{2}t}$

then the survival function of I is

$$\overline{F}(t) = e^{-\lambda_1 t} + \int_{e}^{t} -\lambda_2(t-s) \lambda_1 e^{-\lambda_1 s} ds.$$

Translated into shorthand, the survival function is denoted ${\tt EXP}\,(\, \lambda_{_{\scriptstyle 2}}) \; + \; {\tt EXP}\,(\, \lambda_{_{\scriptstyle 2}}\,) \; .$



This shorthand notation is heuristically apparent. We can visualize a 1 component / 1 spare system with $\mathrm{EXp}(\lambda_A)$ and $\mathrm{Exp}(\lambda_2)$ lives respectively. From component 1 the system has an $\mathrm{EXP}(\lambda_A)$ life to begin with. When component 1 fails, the system has an extra $\mathrm{EXP}(\lambda_2)$ life.

C. MIXTURE OF DISTRIBUTIONS

1. MIX-Notation

In the previous chapter, we formed the sum of independent random lives, which each had weight one, i.e.

$$T = T_1 + T_2.$$

Now consider

$$T = \begin{cases} T_1 & \text{with probability } p_1 \\ \\ T_2 & \text{with probability } p_2 \end{cases}$$

where $p_1 + p_2 = 1$.

Let D_1 and D_2 be the probability distributions of the random variables T_1 and T_2 respectively. The corresponding survival functions are \overline{F}_1 (t) and \overline{F}_2 (t).

Then

$$\overline{F}(t) = p_1 \overline{F}_2(t) + p_2 \overline{F}_2(t)$$
.

In shorthand, the mixture of distributions D_1 and D_2 with respect to the mixing probabilities p_1 and p_2 is denoted



2. <u>Distributive Law</u>

Now let

$$T = T_3 + T^*$$

where

$$T' = \begin{cases} T_1 & \text{with probability p} \\ \\ T_2 & \text{with probability 1-p.} \end{cases}$$

Then

$$T = T_3 + \begin{cases} T_4 \text{ with probability p} \\ T_2 \text{ with probability 1-p.} \end{cases}$$

$$T = \begin{cases} T_3 + T_4 \text{ with probability p} \\ T_3 + T_2 \text{ with probability 1-p.} \end{cases}$$

The distributive law holds due to the fact that the sum of the mixing probabilities for T_1 and T_2 is equal to one.

The survival function of I can be found by convolution:

$$\vec{F}(t) = \vec{F}_3(t) + \int_0^t (p\vec{F}_1(t-s) + (1-p)\vec{F}_2(t-s)) f_3(s) ds.$$

With D_4 , D_2 , D_3 being the probability distributions for T_4 , T_2 , T_3 , the distributive law can be applied to the shorthand notation:

$$D_3 + MIX [pD_1, (1-p)D_2] = MIX [p((D_1 + D_3), (1-p)(D_2 + D_3)].$$



Graphically this can be represented as follows:

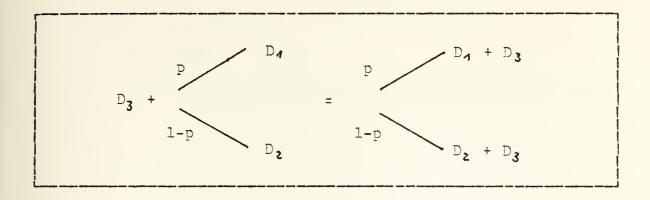


Figure 1: Distributive Property of the MIX-Notation

3. Degeneracy at the Origin

Let

$$P(T=0) = 1.$$

Then the distribution of I is degenerate at zero.

In shorthand notation, such a distribution is called the ZERO-distribution.

Now let $T = T_a + T_o$

where T_a and T_o have probability distributions D_a and ZERO and survival functions \vec{F}_a (t) and \vec{F}_o (t) respectively.

Then

$$\vec{F}(t) = \vec{F}_1(t) + \int_0^t \vec{F}_0(t-s) f_1(s) ds$$
$$= \vec{F}_1(t).$$



The ZERO-distribution doesn't add anything to another distribution, so for instance

$$D_1 + ZERO = D_1$$

$$D_2 + MIX[pD_1, (1-p)ZERO] = MIX[p(D_1 + D_2), (1-p)D_2].$$

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III. APPLYING A RELIABILITY SHORTHAND

After this brief survey over the concept of a reliability shorthand we will now show how the shorthand can be applied. To do so we will use basic structures. Part A of this chapter will give examples whose representation in shorthand requires only basic notation described in Chapter II, Parts A and B, whereas Part B of this chapter will give examples whose representation in shorthand makes use of the MIX-notation and the ZERO-distribution.

A. SUMS OF EXPONENTIALS WITH WEIGHT ONE

1. Simple Series System

A series system is a system which is functioning, when all its components are functioning. A two-component series system can be graphically represented as shown in Fig.2.

Let

 ${\tt T}$: life of the system

 T_4 : life of component 1

T,: life cf component 2

 \vec{F}_1 (t) = survival function of component 1 = $e^{-\lambda_1 t}$



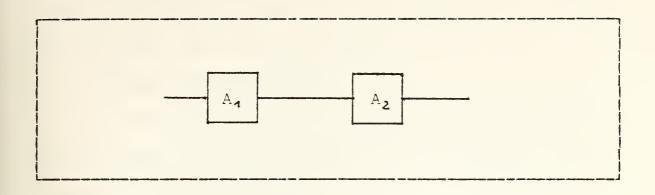


Figure 2: Two-Component Series System

$$\vec{F}_2$$
 (t) = survival function of component 2
= $e^{-\lambda_2 t}$.

Then

$$T = min(T_1, T_2)$$
 $\vec{F}(t) = survival function of the system$

$$= P(min(T_1, T_2) > t)$$

$$= P(T_1 > t, T_2 > t)$$

Assuming independence of the two components

$$\overline{F}(t) = P(T_1 > t) P(T_2 > t)$$

$$= \overline{F}_1(t) \overline{F}_2(t)$$

$$= e^{-\lambda_1 t} e^{-\lambda_2 t}$$

$$= e^{-(\lambda_1 + \lambda_2) t}.$$

The shorthand notation for this system is

EXP
$$(\lambda_0 + \lambda_2)$$
.



This is intuitively apparent, as the system has an exponential survival function with failure rate λ_4 + λ_2 .

2. Simple Parallel System

A parallel system is a system which is functioning, when at least one of its components is functioning. A two-component parallel system can be graphically represented as follows:

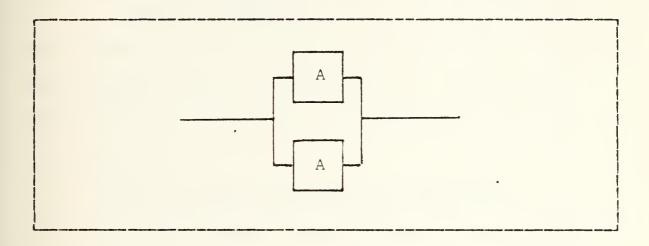


Figure 3: Two-Component Parallel System

Let

$$T_4 \sim EXP(\lambda)$$
, $T_2 \sim EXP(\lambda)$.

Then

$$T = \max (T_1, T_2)$$
 $\overline{F}(t) = P(\max (T_1, T_2) > t)$
 $= 1 - P(\max (T_1, T_2) \le t)$
 $= 1 - P(T_1 \le t, T_2 \le t)$



Assuming independence of the two components,

$$\vec{F}(t) = 1 - P(T_1 \le t) P(T_2 \le t)$$

$$= 1 - F_1(t) F_2(t)$$

$$= 1 - (1 - e^{-\lambda t}) (1 - e^{-\lambda t})$$

$$= 1 - (1 - 2e^{-\lambda t} + e^{-2\lambda t})$$

$$= 2e^{-\lambda t} - e^{-2\lambda t}.$$

The shorthand notation for the system is

$$EXP(2\lambda) + EXP(\lambda)$$
.

This follows intuition as the system has an $EXP(2\lambda)$ life to begin with and when one component fails it has an extra $EXP(\lambda)$ life due to the memoryless property of the exponential distribution.

3. Standby-System with Dissimilar Components

Suppose a system consists of two components, one active and one spare. The active component stays in service until it fails and then immediately is replaced by the spare.

Let the time to failure of the two components be $T_1 \sim \text{EXP}\,(\lambda_1\,) \text{ and } T_2 \sim \text{EXP}\,(\lambda_2\,) \text{ respectively.}$

Then the system time to failure is

$$T = T_1 + T_2$$

and the survival function of the system is

$$\overline{F}(t) = P(T > t)$$



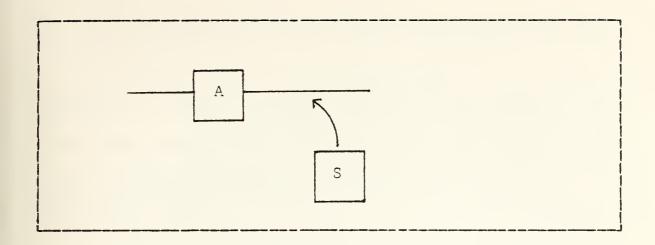


Figure 4: Standby System

$$= \overline{F}_{1}(t) + \int_{0}^{t} \overline{F}_{2}(t-s) f_{1}(s) ds$$

$$= e^{-\lambda_{1}t} + \int_{0}^{t} e^{-\lambda_{2}(t-s)} \lambda_{1} e^{-\lambda_{1}s} ds$$

$$= \frac{\lambda_{1}}{\lambda_{1}-\lambda_{2}} e^{-\lambda_{2}t} - \frac{\lambda_{2}}{\lambda_{1}-\lambda_{2}} e^{-\lambda_{1}t}$$

The shorthand notation for the system's survival function should be obvious. The system has an EXP(λ_4) life from the active component and an additional EXP(λ_2) life from the spare. So the shorthand notation is

$$EXP(\lambda_1) + EXP(\lambda_2)$$
.



B. SUMS OF EXPONENTIALS WITH WEIGHT BETWEEN ZERO AND ONE

The examples given in the previous chapter only involved exponential lives with weight one. Now we will look at some structures, whose survival function has a shorthand notation which includes the MIX-notation and/or the ZERO-distribution.

1. Parallel System with Dissimilar Failure Rates

The notion of a parallel system has been introduced
in Chapter III.A.2. We now look at the case where

$$T_1 \sim EXP(\lambda_1)$$
 and $T_2 \sim EXP(\lambda_2)$.

Then

$$T = \max(T_1, T_2)$$

$$\overline{F}(t) = P(\max(T_1, T_2) > t)$$

$$= 1-P(\max(T_1, T_2) \leq t)$$

$$= 1-P(T_1 \leq t, T_2 \leq t)$$

Assuming independence of the two components

$$\vec{F}(t) = 1 - P(T_1 \leq t) \quad P(T_2 \leq t) \\
= 1 - F_1(t) \quad F_2(t) \\
= 1 - (1 - e^{-\lambda_1 t}) (1 - e^{-\lambda_2 t}) \\
= 1 - (1 - e^{-\lambda_1 t}) (1 - e^{-\lambda_2 t}) \\
= 1 - (1 - e^{-\lambda_1 t}) (1 - e^{-\lambda_2 t}) \\
= e^{-\lambda_1 t} + e^{-\lambda_2 t} - e^{-(\lambda_1 + \lambda_2)t}$$



To find the shorthand notation of the system consider all the ways which lead to the survival of the system:

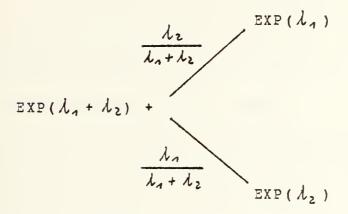
-either both components survive

-or component 1 fails and component 2 survives

-or component 2 fails and component 1 survives.

If one component fails and one survives, in $\frac{\lambda_2}{\lambda_1 + \lambda_2}$ fraction of the cases the survivor will be component 1 and in $\frac{\lambda_1}{\lambda_1 + \lambda_2}$ fraction of the cases it will be component 2.

This can graphically be represented as



Making use of the MIX-notation the shorthand notation then is

$$\text{EXP}(\lambda_1 + \lambda_2) + \text{MIX}[\frac{\lambda_2}{\lambda_1 + \lambda_2} \text{EXP}(\lambda_1), \frac{\lambda_1}{\lambda_1 + \lambda_2} \text{EXP}(\lambda_2)]$$
 and using the distributive property it becomes

$$MIX\left[\frac{\lambda_2}{\lambda_1 + \lambda_2}(EXP(\lambda_1) + EXP(\lambda_1 + \lambda_2))\right],$$

$$\frac{l_1}{l_1+l_2}(\text{EXP}(l_2) + \text{EXP}(l_1+l_2))].$$



As a check to see that this shorthand notation represents the survival function of the system, we derive the survival function from the shorthand notation:

$$\bar{F}(t) = \frac{\lambda_{2}}{\lambda_{n} + \lambda_{2}} \left(e^{-\lambda_{n}t} + \int_{0}^{t} e^{-(\lambda_{n} + \lambda_{2})(t-s)} \lambda_{n} e^{-\lambda_{n}s} ds \right)
+ \frac{\lambda_{n}}{\lambda_{n} + \lambda_{2}} \left(e^{-\lambda_{2}t} + \int_{0}^{t} e^{-(\lambda_{n} + \lambda_{2})(t-s)} \lambda_{2} e^{-\lambda_{2}s} ds \right)
= e^{-\lambda_{n}t} + e^{-\lambda_{2}t} - e^{-(\lambda_{n} + \lambda_{2})t} .$$

This verifies that the shorthand notation indeed represents the system's survival function.

2. Series System with One Spare

Let us now look at a two-component series system, whose components have dissimilar failure rates with one component having a spare:

Component 1 has the constant failure rate \mathcal{L}_1 and component 2 and the spare have the constant failure rate \mathcal{L}_2 . The spare can only replace component 2.

Let

F, (t): the survival function of component 1

 \overline{F}_{2} (t) : the survival function of the standby system component 2 with its spare.



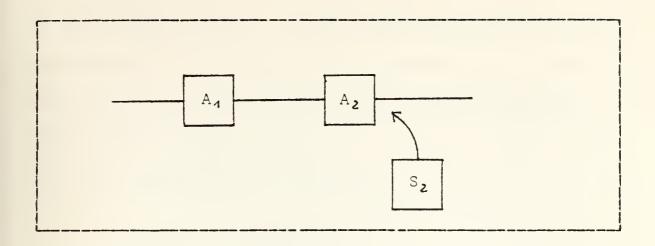


Figure 5: Series System with one Spare

The survival function for a standby system was derived in

Chapter II.B. Therefore

$$\overline{F}_{2}(t) = e^{-\lambda_{2}t} + \int_{0}^{t} e^{-\lambda_{2}(t-s)} \lambda_{2} e^{-\lambda_{2}s} ds$$

$$= e^{-\lambda_{2}t} + \lambda_{2}e^{-\lambda_{2}t} \int_{0}^{t} ds$$

$$= (1 + \lambda_{2}t) e^{-\lambda_{2}t}.$$

Now
$$\vec{F}_1(t) = e^{-\lambda_1 t}$$

Then
$$\overline{F}(t) = \overline{F}_{1}(t) \overline{F}_{2}(t)$$

$$= (1 + \lambda_{2} t) e^{-(\lambda_{1} + \lambda_{2})t}.$$

To translate the survival function into shorthand notation, let us consider the ways in which the system can survive:

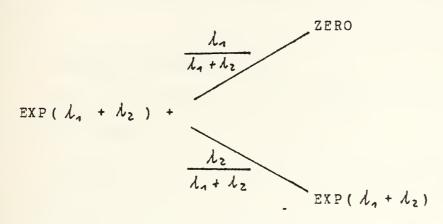
-either both components survive

-or component 2 fails and its spare survives.



If one component fails, in $\frac{\lambda_A}{\lambda_A + \lambda_2}$ fraction of the time it will be component 1, which means that the system will not survive; in $\frac{\lambda_2}{\lambda_A + \lambda_2}$ fraction of the time the failing component will be component 2.

This can graphically be represented as



Using the MIX-notation the survival function then is

$$\begin{aligned} & \text{EXP} \left(\lambda_{1} + \lambda_{2} \right) + \text{MIX} \left[\frac{\lambda_{1}}{\lambda_{1} + \lambda_{2}} \text{ZERO}, \frac{\lambda_{2}}{\lambda_{1} + \lambda_{2}} \text{EXP} \left(\lambda_{1} + \lambda_{2} \right) \right] \\ & = \text{MIX} \left[\frac{\lambda_{1}}{\lambda_{1} + \lambda_{2}} \left(\text{ZERO} + \text{EXP} \left(\lambda_{1} + \lambda_{2} \right) \right), \\ & \frac{\lambda_{2}}{\lambda_{1} + \lambda_{2}} \left(\text{EXP} \left(\lambda_{1} + \lambda_{2} \right) + \text{EXP} \left(\lambda_{1} + \lambda_{2} \right) \right) \right] \\ & = \text{MIX} \left[\frac{\lambda_{1}}{\lambda_{1} + \lambda_{2}} \left(\text{EXP} \left(\lambda_{1} + \lambda_{2} \right) + \text{EXP} \left(\lambda_{1} + \lambda_{2} \right) \right) \right]. \end{aligned}$$

To prove, that the shorthand notation does represent the survival function, we derive the latter from the shorthand:

$$\overline{F}(t) = \frac{\lambda_1}{\lambda_1 + \lambda_2} e^{-(\lambda_1 + \lambda_2)t} + \frac{\lambda_2}{\lambda_1 + \lambda_2} \left(e^{-(\lambda_1 + \lambda_2)t} + \frac{\lambda_2}{\lambda_1 + \lambda_2}\right)$$



$$\int_{0}^{t} e^{-(\lambda_{1}+\lambda_{2})(t-s)} (\lambda_{1}+\lambda_{2}) e^{-(\lambda_{1}+\lambda_{2})s} ds$$

$$= (1+\lambda_{2}t) = -(\lambda_{1}+\lambda_{2})t$$

This is the previously found result and this verifies, that the shorthand notation does represent the system's survival function.

3. Two-out-of-Three System

As a last example in this chapter, we will look at a Two-out-of- Three system.

Consider a three component system, whose components have constant failure rates λ_1 , λ_2 and λ_3 respectively. The system is functioning, as long as two out of three components are functioning (see Fig. 6).

In other words, the system is functioning as long as there is a path through the system .

Alternatively, the system can be visualized as a parallel-series system (compare Fig. 7).

The survival function of the system is

$$\vec{F}(t) = P(T_1 > t \land T_2 > t) + P(T_1 > t \land T_3 > t)
+ P(T_2 > t \land T_3 > t)
- P((T_1 > t \land T_2 > t) \land (T_1 > t \land T_3 > t))
- P((T_1 > t \land T_2 > t) \land (T_2 > t \land T_3 > t))
- P((T_1 > t \land T_3 > t) \land (T_2 > t \land T_3 > t))$$



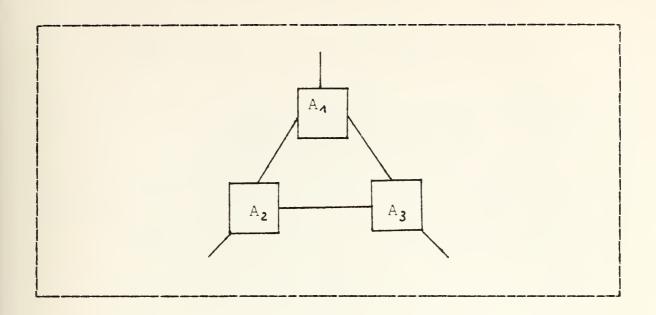


Figure 6: Two-out-of-Three System

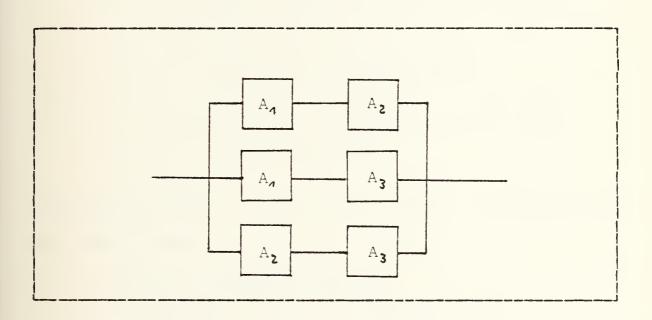


Figure 7: Two-out-of-Three System

+ P((
$$T_1$$
 >t \land T_2 >t) \land (T_4 >t \land T_3 >t)
 \land P(T_2 >t \land T_3 >t)).



Thus

$$\widetilde{F}(t) = P(T_1 > t \land T_2 > t) + P(T_1 > t \land T_3 > t)$$

$$+ P(T_2 > t \land T_3 > t)$$

$$- 3P(T_1 > t \land T_2 > t \land T_3 > t)$$

$$+ P(T_2 > t \land T_2 > t \land T_3 > t)$$

$$+ P(T_2 > t \land T_2 > t \land T_3 > t)$$

Therefore, and assuming independence of the components,

$$\begin{split} \overline{F}(t) &= P(T_{4} > t) P(T_{2} > t) + P(T_{4} > T) P(T_{3} > t) \\ &+ P(T_{2} > t) P(T_{3} > t) \\ &- 3P(T_{4} > t) P(T_{2} > t) P(T_{3} > t) \\ &+ P(T_{4} > t) P(T_{2} > t) P(T_{3} > t) \\ &+ P(T_{4} > t) P(T_{2} > t) P(T_{3} > t) \\ &= P(T_{4} > t) P(T_{2} > t) + P(T_{4} > t) P(T_{3} > t) \\ &+ P(T_{2} > t) P(T_{3} > t) \\ &- 2P(T_{4} > t) P(T_{2} > t) P(T_{3} > t) \\ &= e^{-(\lambda_{4} + \lambda_{3})t} + e^{-(\lambda_{4} + \lambda_{3})t} \\ &- 2e^{-(\lambda_{4} + \lambda_{2} + \lambda_{3})t} \end{split}$$

Now let us consider all the possible ways, in which the system can survive:

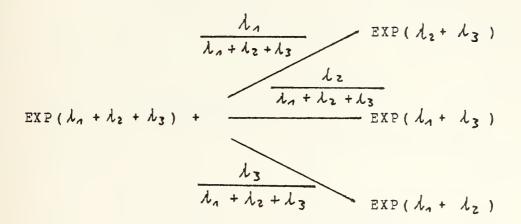
- either all components survive
- or component 1 fails and component 2 and 3
 survive
- or component 2 fails and component 1 and 3 survive



- or component 3 fails and component 1 and 2
survive.

If a component fails and the other two survive, in $\frac{\lambda_i}{\lambda_a + \lambda_z + \lambda_3}$ fraction of the time it will be component i, i = 1,2,3.

This can graphically be represented as



The shorthand notation then is

$$\begin{split} \text{EXP} (\, \lambda_{1} + \lambda_{2} + \lambda_{3}) \, + \, \text{MIX} [\, \frac{\lambda_{1}}{\lambda_{1} + \lambda_{2} + \lambda_{3}} \, \text{EXP} (\, \lambda_{2} + \lambda_{3}) \, , \\ \frac{\lambda_{2}}{\lambda_{1} + \lambda_{2} + \lambda_{3}} \, \text{EXP} (\, \lambda_{1} + \lambda_{3}) \, , \\ \frac{\lambda_{3}}{\lambda_{1} + \lambda_{2} + \lambda_{3}} \, \text{EXP} (\, \lambda_{1} + \lambda_{2}) \,], \end{split}$$

$$= \, \text{MIX} [\, \frac{\lambda_{1}}{\lambda_{1} + \lambda_{2} + \lambda_{3}} \, (\text{EXP} (\, \lambda_{2} + \lambda_{3}) \, + \, \text{EXP} (\, \lambda_{1} + \lambda_{2} + \lambda_{3})) \, , \\ \frac{\lambda_{2}}{\lambda_{1} + \lambda_{2} + \lambda_{3}} \, (\text{EXP} (\, \lambda_{1} + \lambda_{3}) \, + \, \text{EXP} (\, \lambda_{1} + \lambda_{2} + \lambda_{3})) \, , \\ \frac{\lambda_{3}}{\lambda_{2} + \lambda_{3}} \, (\text{EXP} (\, \lambda_{1} + \lambda_{2}) \, + \, \text{EXP} (\, \lambda_{1} + \lambda_{2} + \lambda_{3})) \,]. \end{split}$$



Again, as a check that the shorthand notation represents the survival function, let us derive the survival function from the shorthand notation:

$$F(t) = \frac{\lambda_{1}}{\lambda_{1} + \lambda_{2} + \lambda_{3}} \left[e^{-(\lambda_{1} + \lambda_{2} + \lambda_{3})t} + \int_{e}^{-(\lambda_{1} + \lambda_{2} + \lambda_{3})(t-s)} (\lambda_{2} + \lambda_{3})e^{-(\lambda_{2} + \lambda_{3})s} \right]$$

$$+ \frac{\lambda_{2}}{\lambda_{1} + \lambda_{2} + \lambda_{3}} \left[e^{-(\lambda_{1} + \lambda_{3})t} + \int_{e}^{-(\lambda_{1} + \lambda_{2} + \lambda_{3})(t-s)} (\lambda_{1} + \lambda_{3})e^{-(\lambda_{1} + \lambda_{3})s} \right]$$

$$+ \frac{\lambda_{3}}{\lambda_{1} + \lambda_{2} + \lambda_{3}} \left[e^{-(\lambda_{1} + \lambda_{2} + \lambda_{3})(t-s)} (\lambda_{1} + \lambda_{2})e^{-(\lambda_{1} + \lambda_{2})s} \right]$$

$$+ \int_{e}^{-(\lambda_{1} + \lambda_{2} + \lambda_{3})(t-s)} (\lambda_{1} + \lambda_{2})e^{-(\lambda_{1} + \lambda_{2})s} ds$$

$$= e^{-(\lambda_{1} + \lambda_{2} + \lambda_{3})t} + e^{-(\lambda_{1} + \lambda_{2} + \lambda_{3})t}$$

$$- 2 e^{-(\lambda_{1} + \lambda_{2} + \lambda_{3})t}$$

The result again proves that the shorthand notation indeed represents the survival function of the system.



IV. IMPLEMENTING THE SHORTHAND ON THE II-59
The concept of a reliability shorthand is introduced in the course "Reliability and Weapons System Effectiveness Measurements", OA 4302, at the Naval Postgraduate School, Monterey. Most students taking the course are in the Operations Research (OR) - Curriculum.

The choice of the TI-59 as the computing device, on which the shorthand was to be implemented, was based on the fact, that each student in the OR-Curriculum is issued a TI-59 for use in basic probability and statistics courses. Thus almost every student at the Naval Postgraduate School, who is introduced to the shorthand, is familiar with the TI-59 and has access to such a calculator.

A program, that uses the shorthand notation, times to failure and failure rates as input, should

- calculate the survival probability of basic structures / small systems and
- require moderate computation time.

To achieve these requirements it was decided to incorporate all solutions for the convolution of up to four exponential random variables in the program. The formulas that were used are given in Appendix A.



Two programs are provided in this paper.

Program 1 can be used when all rates are dissimilar or all are the same. It uses the formulas on pages 37 and 38 only.

Program 2 can be used for the general case. It makes use of all the formulas given in Appendix A. The program includes a sorting routine that determines the applicable formula from the entered failure rates.

A user guide to the two programs is provided in Appendix B.



V. SUMMARY

There is a reliability shorthand that denotes the survival function of a system, assuming that the failure rates of all components are constant.

This shorthand can be implemented on the TI-59 handheld calculator. With failure rates, time to failure and shorthand as input the TI-59 calculates the survival probability of the system.

Knowledge of calculus is not necessary to use this method, whereas the standard procedure, finding the survival probability by convolution, requires knowledge of integral calculus.

The choice of the TI-59 as the computing device for the implementation of the shorthand, though, implied limitations; the number of failure rates is limited due to the limited storage capacity of the TI-59, and computing times are comparatively long. The TI-59 can therefore only be used for smaller systems, preferably for the solution of class-room problems.

For the solution of larger problems, the shorthand should be implemented on a state-of-the-art personal



computer using a general algorithm for the convolution of any number of exponential random variables.

APPENDIX A

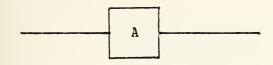
CONVOLUTION FORMULAS

Appendix A contains formulas for the convolution of up to four exponential random variables.

For the two special cases, when all random variables have the same failure rate and all have different failure rates, general formulas for the convolution of any number of exponential random variables are given.

.These formulas are used in the two TI-59 programs provided in Appendix B.



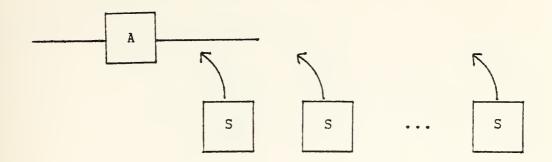


Shorthand: EXP(1)

Survival Function: $\overline{F}(t) = e^{-\lambda t}$

Description:

A single active component with constant failure rate λ .



Shorthand:

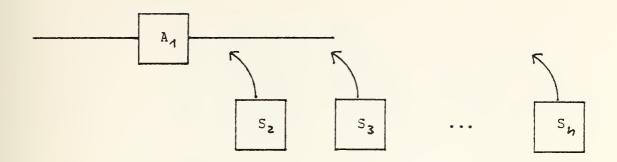
$$EXP(\lambda) + EXP(\lambda) + ... + EXP(\lambda)$$

Survival Function:
$$\overline{F}(t) = \left(\frac{(\lambda t)^0}{0!} + \frac{(\lambda t)^4}{4!} + \cdots + \frac{(\lambda t)^{h-1}}{(h-1)!}\right) e^{-\lambda t}$$

$$= \sum_{i=1}^{n} \frac{(\lambda t)^{i-1}}{(i-1)!} e^{-\lambda t}$$

Description:

A single active component with constant failure rate is supported by n-1 identical spares.



Shorthand:

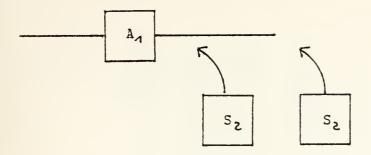
$$EXP(\lambda_1) + EXP(\lambda_2) + \dots + EXP(\lambda_n)$$

Survival Function:
$$\overline{F}(t) = \sum_{i=1}^{n} \left(\frac{\lambda_{i}}{\lambda_{j} - \lambda_{i}} e^{-\lambda_{i} t} \right)$$

Description:

A single active component with constant failure rate is supported by n-1 spares. The active component and the spares have all constant, but dissimilar failure rates.





Shorthand: $EXP(\lambda_1) + EXP(\lambda_2) + EXP(\lambda_2)$

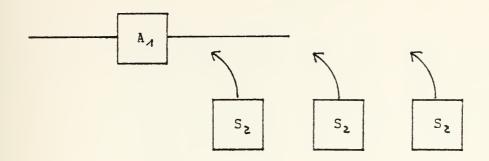
Survival Function:
$$\overline{F}(t) = Ae^{-\lambda_A t} + (B + Ct) e^{-\lambda_2 t}$$

where $A = \frac{\lambda_z^2}{(\lambda_2 - \lambda_A)^2}$
 $B = 1 - A$
 $C = \frac{\lambda_A \lambda_2}{\lambda_A - \lambda_2}$

Description:

A single active component with constant failure rate λ_4 is supported by two spares with identical constant failure rate λ_2 .





Shorthand: $EXP(\lambda_2) + EXP(\lambda_2)$

$$EXP(\lambda_2) + EXP(\lambda_2) + EXP(\lambda_2) + EXP(\lambda_2)$$

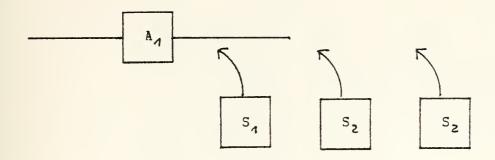
Survival Function:
$$\overline{F}(t) = Ae^{-\lambda_n t} + (B + Ct + Dt^2)e^{-\lambda_2 t}$$

where $A = \frac{\lambda_2^3}{(\lambda_2 - \lambda_n)^3}$
 $B = 1 - A$
 $C = \lambda_2 - \frac{\lambda_2^3}{(\lambda_n - \lambda_2)^2}$
 $D = \frac{\lambda_n \lambda_2^2}{2(\lambda_n - \lambda_2)}$

Description:

A single active component with constant failure rate λ_2 is supported by three spares with identical constant failure rate λ_2 .





Shorthand:
$$EXP(\lambda_1) + EXP(\lambda_1) + EXP(\lambda_2) + EXP(\lambda_2)$$

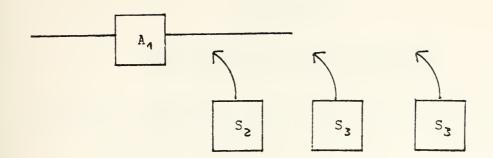
Survival Function:
$$F(t) = (A + Bt)e^{-\lambda_n t} + (C + Dt)e^{-\lambda_2 t}$$

where $A = \frac{\lambda_2^3 - 3\lambda_2^2 \lambda_n}{(\lambda_2 - \lambda_n)^3}$
 $B = \frac{\lambda_n \lambda_2^2}{(\lambda_2 - \lambda_n)^2}$
 $C = 1 - A$
 $D = \frac{\lambda_n^2 \lambda_2}{(\lambda_n - \lambda_2)^2}$

Description:

A single active component with constant failure rate λ_A is supported by one identical spare and two spares with dissimilar, constant failure rate λ_A .





$$EXP(\lambda_1) + EXP(\lambda_2) + EXP(\lambda_3) + EXP(\lambda_3)$$

Survival Function:
$$\overline{F}(t) = Ae^{-\lambda_0 t} + Be^{-\lambda_2 t} + (C + Dt)e^{-\lambda_3 t}$$

where
$$A = \frac{\lambda_2 \lambda_3^2}{(\lambda_2 - \lambda_4)(\lambda_3 - \lambda_A)^2}$$

$$B = \frac{\lambda_1 \lambda_3^2}{(\lambda_4 - \lambda_2)(\lambda_3 - \lambda_2)^2}$$

$$C = \frac{\lambda_1 \lambda_2}{(\lambda_4 - \lambda_3)(\lambda_2 - \lambda_3)} + \frac{\lambda_1 \lambda_2 \lambda_3}{(\lambda_4 - \lambda_2)} \left(\frac{1}{(\lambda_4 - \lambda_3)^2} - \frac{1}{(\lambda_2 - \lambda_3)^2}\right)$$

$$D = \frac{\lambda_1 \lambda_2 \lambda_3}{(\lambda_4 - \lambda_3)(\lambda_2 - \lambda_3)}$$

Description:

A single active component with constant failure rate λ_2 has three spares. One spare has constant failure rate λ_2 , two spares are identical with constant failure rate λ_3 .



APPENDIX B

USER GUIDE TO TI-59 PROGRAMS

Appendix B contains a user guide to two TI-59 programs, which use reliability shorthand and failure rates as input to compute the survival probability of a system.

PROGRAM 1 is designed for the two special cases where the reliability shorthand is of the form

$$EXP(\lambda) + EXP(\lambda) + ... + EXP(\lambda)$$

or

$$EXP(\lambda_1) + EXP(\lambda_2) + ... + EXP(\lambda_h)$$
.

In the first case the number of terms is not limited, whereas in the second case the number of terms is limited to 40 due to limited storage capacity of the TI-59. In this case the number of terms can be increased to 70 by entering 9 in the display and pressing 2nd Op 17.

PROGRAM 2 is designed to solve problems of the kind, that were introduced in Chapter III.B. . Due to limited memory of the TI-59 the number of exponential terms under one weight in shorthand notation is limited to four.



All results will be printed, if the TI-59 is connected to a TI PC-100A or TI PC-100C printer.

All complete will be printed, in the TI-59 in descend

PROGRAM 1 : Procedure

- Use any library module. Read in program 1 (side 1 of the magnetic card)
- 2. Press 2nd C' to initialize.
- 3. Enter n, the number of exponential terms to be convolved, in the display and press A.
- 4. Enter time t and press B.
- 5. Enter λ_i and press C . When all failure rates are the same, enter λ only once.
- 6. a) To find the survival probability of the system,
 when all failure rates are the same, press 2nd A'.
 - b) To find the survival probability of the system,
 when all failure rates are dissimilar, press 2nd B.



PROGRAM 1 : Sample Problems

1. Find the survival probability of a parallel system

(compare Chapter III.A. 2)

a)
$$l = .3$$
, $t = 7$, $n = 2$

b) Shorthand notation:

$$EXP(.6) + EXP(.3)$$

c)	Enter	Comment	Press	Display	
		Initialize	C *	0	
	2	n	A	0	
	7	t	В	7	
	. 6	2 L	С	. 3	
	. 3	l	С	. 3	
		F(t)	В	.22991727	797

calculation takes 13 seconds

2. Find the survival probability of a standby-system with dissimilar components (compare Chapter III.A.3) .

a)
$$l_1 = .4$$
, $l_2 = .5$, $t = 6$, $n = 2$

b) Shorthand notation:

$$EXP(.4) + EXP(.5)$$

c)	Enter	Comment	Press	Display
		Initialize	C *	Э
	2	n	A	0
	6	t	В	6
	- 4	da	С	. 4
	. 5	λ_z	С	. 5
		F (t)	В •	. 254441493

calculation takes 13 seconds



- 3. Find the survival probability of a standby-system with one active component and four similar spares.
 - a) l = .3 , t = 7 , n = 5
 - b) Shorthand notation:

$$EXP(.3) + EXP(.3) + EXP(.3) + EXP(.3) + EXP(.3)$$

c)	Enter	Comment	Press	Display	
		Initialize	C 1	0	
	5	n	A	0	
	7	t	В	7	
	. 3	l	С	. 3	
	•	F (t)	A *	.937873884	3

calculation takes 9 seconds



- PROGRAM 2 : Procedure
- CASE I: To find the convolution of up to four exponential random variables.
- 1. Use any library module.
 Re-Partition (enter 2 in the display, press 2nd Op 17).

Read in all four sides of the magnetic card.

- 2. Press 2nd C' to initialize.
- 3. Enter n, the number of exponential terms to be convolved, in the display and press A.
- 4. Enter time t and press B.
- 5. Enter λ_i and press C (n entries) .
 - REMARK: Failure rates, which appear only once in the expression, have to be entered before failure rates, that appear several times.
- 6. To find the survival probability of the system press E.



PROGRAM 2, CASE I : Sample Problems

(1) Shorthand notation

$$EXP(\lambda_4) + EXP(\lambda_2) + EXP(\lambda_2)$$

Sample values: λ_4 = .3 , λ_2 = .4 , t = 7

Procedure :

Enter	Comment	Press	Display
	Initialize	C 1	0
3	n	A	Э
7	t	В	7
. 3	da	С	. 3
. 4	d 2	C	. 4
. 4	Lz	С	. 4
	F(t)	E	.5363473866

calculation takes 14 seconds

(2) Shorthand notation

 $\text{EXP}(\lambda_1) + \text{EXP}(\lambda_2) + \text{EXP}(\lambda_2) + \text{EXP}(\lambda_2)$ Sample values: $\lambda_1 = .2$, $\lambda_2 = .4$, t = 3Procedure:

Enter	Comment	Press	Display	
	Initialize	21	0	
4	n	A	9	
3	t	В	3	
. 2	do	С	. 2	
. 4	λz	С	- 4	
. 4	λz	С	- 4	
. 4	λz	C	- 4	
	F (+)	Ξ	.98097460	99

calculation takes 20 seconds

(3) Shorthand notation

EXP(λ_1) + EXP(λ_2) + EXP(λ_2) + EXP(λ_2)

Sample values: λ_1 = .4 , λ_2 = .3 , t = 5

Procedure:

Enter	Comment	Press	Display
	Initialize	C 1	0
4	n	A	0
5	t	Б	5
. 4	d,	С	. 4
. 4	1.	С	. 4
. 3	λz	С	. 3
. 3	Lz	С	. 3
•	F(t)	E	.9029040721
•	F(t)	E	.9029040721

calculation takes 20 seconds

(4) Shorthand notation

Procedure:

Enter	Comment	Press	Display	
	Initialize	C *	Э	
4	n	A	0	
10	t	В	10	
. 1	da	C	. 1	
. 3	12	C	.3	
.5	λ_3	С	• 5	
. 5	13	C	. 5	
	F (t)	Ε	.7312684	703

calculation takes 25 seconds

PROGRAM 2 : Procedure

CASE II: to solve problems of the kind, that were introduced in Chapter III.B. .

- 1. Derive the system's shorthand notation. Find either the
 - graphical representation or
 - the MIX-notation .
- 2. Use any library module.

Re-Partition (enter 2 in the display, press 2nd Op 17).

Read in all four sides of the magnetic card.

- 3. Press 2nd C' to initialize.
- 4. Enter time t and press B.
- 5. Repeat the following steps for each path of the graphical representation, i.e. for each convolution in the MIX-notation.
 - a) Enter n, the number of exponential terms to be convolved, in the display and press A.
 - b) Enter λ_i and press C.

REMARK: Failure rates, which appear only once in the expression, have to be entered before failure rates, that appear several times.

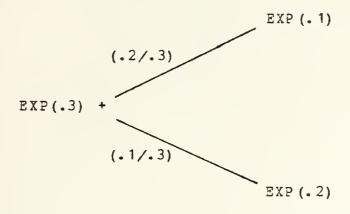
- c) Enter p;, the weight in the ith path, and press D.
- d) To find the part of the system's survival probability, that is contributed by the ith path, press E.



6. To find the survival probability of the system press 2nd E.

PROGRAM 2, CASE II : Sample Problems

- 1. Find the survival probability of a parallel system with dissimilar failure rates (compare Chapter III.B.1).
 - a) $l_1 = .1$, $l_2 = .2$, t = 2
 - b) Shorthand notation

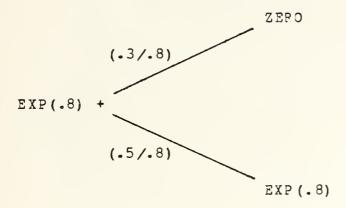


$$F(t) = MIX[(.2/.3) (EXP(.1) + EXP(.3), (.1/.3) (EXP(.2) + EXP(.3)].$$

Procedure :

Enter	Comment	Press	Display
	Initialize	C *	0
2	ŧ	В	2
2	n_A	A	О
.1	1,	С	. 1
•3	$l_n + l_2$	C	. 3
(.2/.3)	Pa	D	.6666666667
		E	.635793541
2	n s	A	Э
. 2	dz	С	• 2
.3	$l_1 + l_2$	C	. 3
(.1/.3)	P ₂	D	.333333333
		E	.304445622
	F (t)	E *	.940239163

- 2. Find the survival probability of a series system with one spare as introduced in Chapter III.B.2.
 - a) $l_1 = .3$, $l_2 = .5$, t = 7
 - b) Shorthand notation



$$\vec{F}(t) = MIX[(.3/.8) (EXP(.8), (.5/.8) (EXP(.8) + EXP(.8)].$$

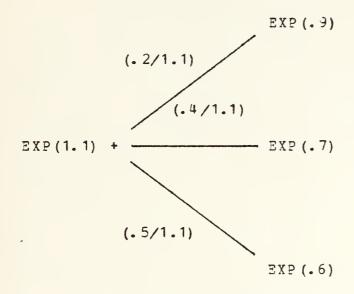
Procedure:

Enter	Comment	Press	Display
	Initialize	C 1	0
7 .	ŧ	В	7
1	n ₁	A	0
. 8	$l_1 + l_2$	С	. 8
(.3/.8)	P ₁	D	.375
		E	.0013866989
2	n 2	A	0
. 8	$d_1 + d_2$	С	.8
. 8	$l_1 + l_2$	С	. 8
(.5/.8)	Pz	D	.625
		Ē	.0152536878
	F(t)	E *	.0166403867

3. Find the survival probability of a Two-out-of-Three System as introduced in Chapter III.B.3.

a)
$$l_1 = .2$$
 , $l_2 = .4$, $l_3 = .5$, $t = 9$

b) Shorthand notation



$$\vec{F}(t) = MIX[(.2/1.1) (EXP(.9) + EXP(1.1)),$$

$$(.4/1.1) (EXP(.7) + EXP(1.1)),$$

$$(.5/1.1) (EXP(.6) + EXP(1.1))].$$

C)

Procedure :

Enter	Comment	Press	Display
	Initialize	C *	0
9	t	В	9
2	n ₁	A	0
1.1	$l_1 + l_2 + l_3$	С	1.1
• 9	$d_2 + d_3$	С	• 9
(.2/1.1)	P ₁	D	.1818181818
		E	.0002624871
2	n s	A	. 0
1.1	$\lambda_1 + \lambda_2 + \lambda_3$	С	1.1
.7	$\lambda_1 + \lambda_3$	С	.7
(.4/1.1)	P _z	D	.3636363636
		Ξ	.0018043754
2	n ₃	A	0
1.1	$l_1 + l_2 + l_3$	С	1.1
.6	$d_1 + d_2$	С	. 6
(.5/1.1)	p ₃	D	.4545454545
		Ε	.0044892129
	F (t)	E •	.0065560755

COMPUTER LISTINGS

PROGRAM 1

0123456789012345678901234567890123456789 000000000001111111111222234567890123456786	SL RSTL 00 1 1 TOP 20 TO 10 TO 10 SL 88 BT 0 16 CT 0 P 2 OT 0 1 TO 10 TO 10 BT 0 NB CT 0 P 2 OT 10 TO 10 ST 1 NB BT 0 NB CT 0 P 2 OT 10 ST 1 NB BT 0 NB CT 0 P 2 OT 10 ST 1 NB BT 0 NB CT 0 P 2 OT 10 ST 1 NB BT 0 NB CT 0 P 2 OT 10 ST 1 NB BT 0 NB CT 0 P 2 OT 10 ST 1 NB BT 0 NB CT 0 P 2 OT 10 ST 1 NB BT 0 NB CT 0 P 2 OT 1 NB CT 0 P 2 OT 1 NB BT 0 NB CT 0 P 2 OT 1 NB CT 0 P 2	0 = 2045570000000000000000000000000000000000	L 10 14 02 040505050616 190 12 15 14 M8 9 1 16 03 CO 1 CO	0:2345478901234547890-2354189012345478 88888889999999990-2351541-21121	15141001 121 VX RS TS1 1818 06 06 06 18 18 86 P 07 08 18 18 18 18 18 18 18 18 18 18 18 18 18
039	28 28 91 R/S	079	25 25	118	06 06 57 EQ



010101010101010101010124444400000000000	3703702709618HND6 *8 XD6L0TL6 19 GM 60NL PR XR 1 FD GSLTP1 CO = /R1C0: L5 = GM 60NL GSLTP1 R XR XR 1 FD GSLTP1 R XR XR 19 GSLTP1 R XR X			\$	M*81 L1 VX L6 M8LOTL8 GR 808LRRSTSL8L8NGCOVXCO=NNXC4#U4CO**CO+49#E/P2F08/BMR/BMC4F
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PROGRAM 2



121 76 LBL 161 69 CP 201 122 38 SIN 162 26 26 202 123 29 CP 163 61 GTD 203 124 73 RC* 164 38 SIN 204 125 06 06 165 76 LBL 205 126 67 EQ 166 36 PGM 206 127 36 PGM 167 73 RC* 207 128 73 RC* 168 08 08 208 129 08 08 169 94 +/- 209 130 32 X‡T 170 65 × 210 131 73 RC* 171 43 RCL 211 132 06 06 172 01 01 212 133 22 INV 173 95 = 213 134 67 EQ 174 22 INV 214 135 30 TAN 175 23 LNX 215 136 69 GP 176 65 × 216 137 26 26 177 43 RCL 216 137 26 26 177 20 1 01 138 61 GTD 178 16 16 137 26 26 177 22 11 138 38 SIN 179 95 = 219 140 76 LBL 180 44 SUM 220 141 30 TAN 181 18 18 221 142 49 PRD 182 43 RCL 223 144 75 - 184 32 X‡T 224 145 73 RC* 185 43 RCL 223 146 08 08 186 08 08 226 147 95 = 187 75 - 227 148 35 1/X 188 09 9 228 149 49 PRD 189 95 = 229 150 16 16 190 67 EQ 230 151 43 RCL 191 37 P/R 231 154 43 RCL 191 37 P/R 231 154 43 RCL 194 61 GTD 234 155 06 06 195 09 CDS 235 156 75 - 196 76 LBL 235 157 09 9 197 37 P/R 237 158 95 = 198 71 SBR 238 159 77 GE 199 47 CMS	18 L7 TM9NLVL1
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0-404567860-4064567880-406456788688888888888888888888888888888888	\$\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\	0-20454740000-40045676000-20045676000-200456	+ < < L2	0-2345474444444444444444444444444444444444	XC1 VX
394 395 396 397 398 399	12 12	435 435 437 433 443 444	95 = 65 %	4456789 4474444	54) 65 X 53 K 43 ROL 12 12 94 +/-



53142354817183318335 64052294817498331839553174163 64123456789012345678901234567 6488889912345678900000000000011234567 641444444444444444455555555555555555555	RCL 01 :) : INV : LMX ; = : SUM	104531423523553334535555555555555555555555	\$686468688877777777778888888888888888888	.1 94 +/- .2 65 X .3 43 RCL
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01234567890+234567890+234567890+234567890+234567890+234567890+344567890+344567890+344678990+344678990+344678990+344678990+344678990-344678990-344678990-344678990-344678990-344678990-344678990-344678990-344678990-344678990-344678990-344678990-344678990-3446789900-3446789900-3446789900-3446789900-3446789900-3446789900-3446789900-3446789900-3446789900-34467899000-34467899000-3446789900000000000000000000000000000000000	1 L2 L0 L1	0123456789012345678901234567890123456789012345678789012345678789012345678787878787878787878787878787878787878	2 L1	012345678901234567890123456789 888888888999999999999999999999999 888888	FFD00 007L1TF1 F2CF3DF41LCL00 D LDL00 L200 S N X F 0 EF 0 R F 0 X B X C 1 E R 60002731271672873974016874 4 16161744161614162
639	43 RCL	679 679	66 SIF	718	22 INV 61 GTO



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BIBLIOGRAPHY

Barlow, R. E. and Proschan, F., Statistical Theory of Reliability and Life Testing, Silver Spring, 1981.

Esary, J.D., Course Notes and Problem Sets for OA 4302, Naval Postgraduate School, Monterey, California, 1982.

Repicky, John J., An Introduction to a Reliability Shorthand, Master's Thesis, Naval Postgraduate School, Monterey, California, 1981.



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